



# Acknowledgements

Professor Peter Sasieni

Professor Jack Cuzick

# Outline

- ▶ Clinical trials
- ▶ Model and hypotheses
- ▶ Test statistics
- ▶ Test performances
- ▶ Summary

# Clinical trials

One or more **active** arms against a **control** arm (placebo or a standard treatment).

# Clinical trials

One or more **active** arms against a **control** arm (placebo or a standard treatment).

Well-developed methodologies are available for comparing one active vs control.

# Clinical trials

One or more **active** arms against a **control** arm (placebo or a standard treatment).

Well-developed methodologies are available for comparing one active vs control.

Increasing demand for multi-arm trials, due to cost effectiveness (and sometimes benefits of having fewer patients on control).

# Clinical trials

One or more **active** arms against a **control** arm (placebo or a standard treatment).

Well-developed methodologies are available for comparing one active vs control.

Increasing demand for multi-arm trials, due to cost effectiveness (and sometimes benefits of having fewer patients on control).

Less methodology for comparing two or more active vs control.

## Model and Notation

Let

$$Y_{ij} \sim N(\mu_i, 1),$$

where  $Y_{ij}$  is the scaled response from patient  $j$  on treatment  $i = 0, 1, 2$ ,  $j = 1, \dots, n_i$ .

## Model and Notation

Let

$$Y_{ij} \sim N(\mu_i, 1),$$

where  $Y_{ij}$  is the scaled response from patient  $j$  on treatment  $i = 0, 1, 2$ ,  $j = 1, \dots, n_i$ .

Let

$$Z_i = \frac{\sqrt{N}(\bar{Y}_i - \bar{Y}_0)}{\sigma}, \quad i = 1, 2,$$

where  $N = n_0 + n_1 + n_2$ ,  $\sigma^2 = \{1 - \delta\}/\{\delta(1 - 2\delta)\}$  and  $\delta = n_1/N = n_2/N$ .

## Model and Notation

Then

$$\begin{bmatrix} Z_1 \\ Z_2 \end{bmatrix} \sim N \left( \begin{bmatrix} \Delta_1 \\ \Delta_2 \end{bmatrix}, \begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix} \right),$$

where

$$\Delta_i = \frac{\sqrt{N}(\mu_i - \mu_0)}{\sigma}$$

and  $\rho = \text{Cov}(Z_i, Z_j) = \delta / \{1 - \delta\}$ ,  $i \neq j$ .

# Hypotheses

For comparing **three (or more)** arms:

$H_0 : \Delta_1 \leq 0$  and  $\Delta_2 \leq 0$  vs  $H_1 : \Delta_i > 0$  for at least one  $i = 1, 2$ .

# Hypotheses

For comparing **three (or more)** arms:

$H_0 : \Delta_1 \leq 0$  and  $\Delta_2 \leq 0$  vs  $H_1 : \Delta_i > 0$  for at least one  $i = 1, 2$ .

**Note:**  $H_0$  is a **composite** hypothesis.

# Hypotheses

For comparing **three (or more)** arms:

$H_0 : \Delta_1 \leq 0$  and  $\Delta_2 \leq 0$  vs  $H_1 : \Delta_i > 0$  for at least one  $i = 1, 2$ .

**Note:**  $H_0$  is a **composite** hypothesis.

Whereas with two arms, a  $Z$ -test is uniformly most powerful (UMP), with three or more arms **no** UMP test is known.

## Test statistics

We consider the following test statistics, where  $A^+ = \max(0, A)$  and  $X_1 = \max(Z_1, Z_2)$  and  $X_2 = \min(Z_1, Z_2)$ :

1.  $T_\infty = X_1^+$ ;

## Test statistics

We consider the following test statistics, where  $A^+ = \max(0, A)$  and  $X_1 = \max(Z_1, Z_2)$  and  $X_2 = \min(Z_1, Z_2)$ :

1.  $T_\infty = X_1^+$ ;
2.  $S_1^0 = Z_1 + Z_2 = X_1 + X_2$ ;

## Test statistics

We consider the following test statistics, where  $A^+ = \max(0, A)$  and  $X_1 = \max(Z_1, Z_2)$  and  $X_2 = \min(Z_1, Z_2)$ :

1.  $T_\infty = X_1^+$ ;
2.  $S_1^0 = Z_1 + Z_2 = X_1 + X_2$ ;
3.  $T_1 = X_1^+ + \frac{(X_2 - \rho X_1)^+}{\sqrt{1 - \rho^2}}$ ;

## Test statistics

We consider the following test statistics, where  $A^+ = \max(0, A)$  and  $X_1 = \max(Z_1, Z_2)$  and  $X_2 = \min(Z_1, Z_2)$ :

1.  $T_\infty = X_1^+$ ;
2.  $S_1^0 = Z_1 + Z_2 = X_1 + X_2$ ;
3.  $T_1 = X_1^+ + \frac{(X_2 - \rho X_1)^+}{\sqrt{1 - \rho^2}}$ ;
4.  $T_2 = \sqrt{X_1^{+2} + \frac{(X_2 - \rho X_1)^{+2}}{1 - \rho^2}}$ .

# Likelihood Ratio Test

$T_2^2$  is the likelihood ratio test statistic given by

$$T_2^2 = 2 \ln \left\{ \frac{L(\mathbf{y}; \hat{\boldsymbol{\mu}})}{L(\mathbf{y}; \tilde{\boldsymbol{\mu}})} \right\},$$

where  $\hat{\boldsymbol{\mu}}$  is the unrestricted maximum likelihood estimator of  $\boldsymbol{\mu}$  under  $H_0 \cup H_1$ ;

and  $\tilde{\boldsymbol{\mu}}$  is the restricted maximum likelihood estimator of  $\boldsymbol{\mu}$  under  $H_0$ .

# Likelihood Ratio Test

$T_2^2$  is the likelihood ratio test statistic given by

$$T_2^2 = 2 \ln \left\{ \frac{L(\mathbf{y}; \hat{\boldsymbol{\mu}})}{L(\mathbf{y}; \tilde{\boldsymbol{\mu}})} \right\},$$

where  $\hat{\boldsymbol{\mu}}$  is the unrestricted maximum likelihood estimator of  $\boldsymbol{\mu}$  under  $H_0 \cup H_1$ ;

and  $\tilde{\boldsymbol{\mu}}$  is the restricted maximum likelihood estimator of  $\boldsymbol{\mu}$  under  $H_0$ .

Restricted model under  $H_0$ ,  $\mu_i \leq \mu_0$ ,  $i = 1, 2$ , defines a **simple tree order**.

# Pool Adjacent Violators Algorithm (PAVA)

Assume  $\hat{\mu}_1 > \hat{\mu}_2$  (other order follows by symmetry).

# Pool Adjacent Violators Algorithm (PAVA)

Assume  $\hat{\mu}_1 > \hat{\mu}_2$  (other order follows by symmetry).

Start with  $\tilde{\mu}_i = \hat{\mu}_i, \forall i = 0, 1, 2 \Rightarrow T_2 = 0$ .

# Pool Adjacent Violators Algorithm (PAVA)

Assume  $\hat{\mu}_1 > \hat{\mu}_2$  (other order follows by symmetry).

Start with  $\tilde{\mu}_i = \hat{\mu}_i, \forall i = 0, 1, 2 \Rightarrow T_2 = 0$ .

If  $\hat{\mu}_1 > \hat{\mu}_0$ ,

▶ let  $\tilde{\mu}_0 = \tilde{\mu}_1 = \frac{n_0\hat{\mu}_0 + n_1\hat{\mu}_1}{n_0 + n_1} \Rightarrow T_2 = X_1$ .

# Pool Adjacent Violators Algorithm (PAVA)

Assume  $\hat{\mu}_1 > \hat{\mu}_2$  (other order follows by symmetry).

Start with  $\tilde{\mu}_i = \hat{\mu}_i, \forall i = 0, 1, 2 \Rightarrow T_2 = 0$ .

If  $\hat{\mu}_1 > \hat{\mu}_0$ ,

▶ let  $\tilde{\mu}_0 = \tilde{\mu}_1 = \frac{n_0\hat{\mu}_0 + n_1\hat{\mu}_1}{n_0 + n_1} \Rightarrow T_2 = X_1$ .

▶ If  $\hat{\mu}_2 > \tilde{\mu}_0$ ,

▶ let

$$\tilde{\mu}_i = \frac{n_0\hat{\mu}_0 + n_1\hat{\mu}_1 + n_2\hat{\mu}_2}{n_0 + n_1 + n_2}, \forall i = 0, 1, 2 \Rightarrow T_2 = \sqrt{X_1^2 + \frac{(X_2 - \rho X_1)^2}{1 - \rho^2}}.$$

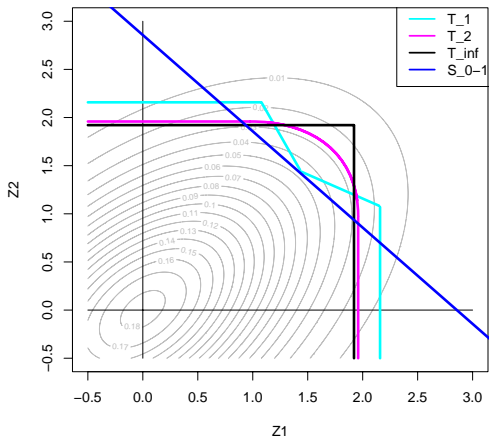
# Null distributions

The null distributions are approximated by Monte Carlo simulation of 1 million draws from the joint null distribution of  $Z_1$  and  $Z_2$ .

Table: Critical values for rejecting  $H_0$ .

Test	5%	2.5%	1%
$T_1$	2.1537	2.5289	2.9839
$T_2$	1.9545	2.2579	2.6120
$T_\infty$	1.9163	2.2121	2.5577
$S_1^0$	2.8488	3.3948	4.0291

### Rejection boundary



## Test performances

We compare the different tests in terms of **power** and the **expected loss** by adding various values of  $\Delta_1$  and  $\Delta_2$  to 100,000 simulations from the joint null distribution of  $Z_1$  and  $Z_2$ .

## Test performances

We compare the different tests in terms of **power** and the **expected loss** by adding various values of  $\Delta_1$  and  $\Delta_2$  to 100,000 simulations from the joint null distribution of  $Z_1$  and  $Z_2$ .

Power = the proportion of simulations which correctly rejected  $H_0$ .

## Test performances

We compare the different tests in terms of **power** and the **expected loss** by adding various values of  $\Delta_1$  and  $\Delta_2$  to 100,000 simulations from the joint null distribution of  $Z_1$  and  $Z_2$ .

Power = the proportion of simulations which correctly rejected  $H_0$ .

Incorrect selection:

correctly **reject  $H_0$**  but select the **wrong** experimental arm.

## Test performances

We compare the different tests in terms of **power** and the **expected loss** by adding various values of  $\Delta_1$  and  $\Delta_2$  to 100,000 simulations from the joint null distribution of  $Z_1$  and  $Z_2$ .

Power = the proportion of simulations which correctly rejected  $H_0$ .

Incorrect selection:

correctly **reject  $H_0$**  but select the **wrong** experimental arm.

Expected Loss = the mean Loss over simulations, where

$$\text{Loss} = \begin{cases} 0 & \text{if we select treatment 1;} \\ \Delta_1 - \Delta_2 & \text{if we select treatment 2;} \\ \Delta_1 & \text{if we fail to reject } H_0. \end{cases}$$

## Test performances

We compare the different tests in terms of **power** and the **expected loss** by adding various values of  $\Delta_1$  and  $\Delta_2$  to 100,000 simulations from the joint null distribution of  $Z_1$  and  $Z_2$ .

Power = the proportion of simulations which correctly rejected  $H_0$ .

Incorrect selection:

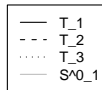
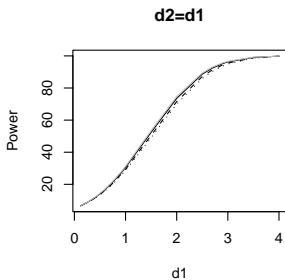
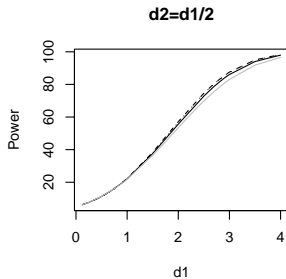
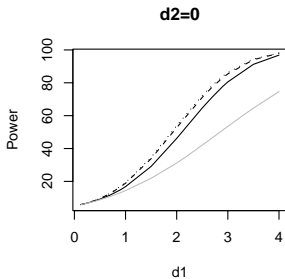
correctly **reject  $H_0$**  but select the **wrong** experimental arm.

Expected Loss = the mean Loss over simulations, where

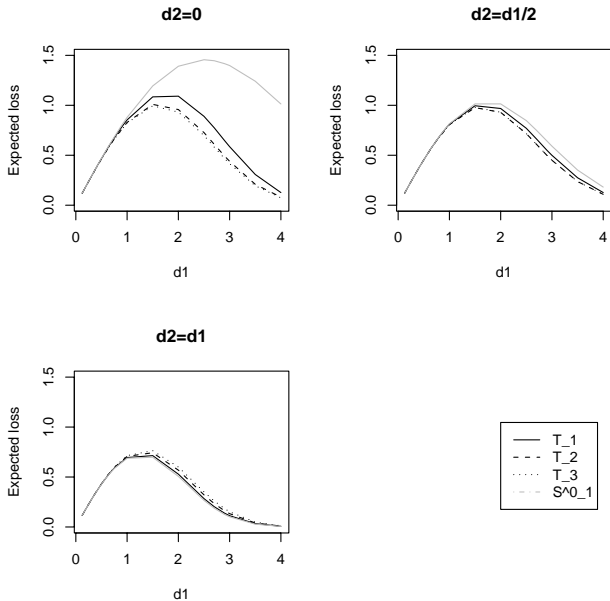
$$\text{Loss} = \begin{cases} 0 & \text{if we select treatment 1;} \\ \Delta_1 - \Delta_2 & \text{if we select treatment 2;} \\ \Delta_1 & \text{if we fail to reject } H_0. \end{cases}$$

We consider correlation  $\rho = 0.5$ , i.e. assuming equal numbers of patients in each arm.

# Power at 5%



# Expected loss at 5%



# Summary

- ▶ No test is uniformly better than the others

# Summary

- ▶ No test is uniformly better than the others
- ▶  $T_\infty$  is best when only one treatment is better than control

# Summary

- ▶ No test is uniformly better than the others
- ▶  $T_\infty$  is best when only one treatment is better than control
- ▶  $T_1$  is better when treatments are similar

# Summary

- ▶ No test is uniformly better than the others
- ▶  $T_\infty$  is best when only one treatment is better than control
- ▶  $T_1$  is better when treatments are similar
- ▶  $T_2$  is more consistent with the best one in each of the above cases and is best in the middle

# Summary

- ▶ No test is uniformly better than the others
- ▶  $T_\infty$  is best when only one treatment is better than control
- ▶  $T_1$  is better when treatments are similar
- ▶  $T_2$  is more consistent with the best one in each of the above cases and is best in the middle
- ▶ Overall, there is not much difference between  $T_2$  and  $T_\infty$

# Summary

- ▶ No test is uniformly better than the others
- ▶  $T_\infty$  is best when only one treatment is better than control
- ▶  $T_1$  is better when treatments are similar
- ▶  $T_2$  is more consistent with the best one in each of the above cases and is best in the middle
- ▶ Overall, there is not much difference between  $T_2$  and  $T_\infty$
- ▶  $T_2$  has rarely (if ever) been used, but is at least as good as  $T_\infty$

# Summary

- ▶ No test is uniformly better than the others
- ▶  $T_\infty$  is best when only one treatment is better than control
- ▶  $T_1$  is better when treatments are similar
- ▶  $T_2$  is more consistent with the best one in each of the above cases and is best in the middle
- ▶ Overall, there is not much difference between  $T_2$  and  $T_\infty$
- ▶  $T_2$  has rarely (if ever) been used, but is at least as good as  $T_\infty$

# Summary

- ▶ No test is uniformly better than the others
- ▶  $T_\infty$  is best when only one treatment is better than control
- ▶  $T_1$  is better when treatments are similar
- ▶  $T_2$  is more consistent with the best one in each of the above cases and is best in the middle
- ▶ Overall, there is not much difference between  $T_2$  and  $T_\infty$
- ▶  $T_2$  has rarely (if ever) been used, but is at least as good as  $T_\infty$

The results can be extended to more arms.

# References

Barlow, Bartholomew, Bremner and Brunk (1972) *Statistical Inference Under Order Restrictions*. Wiley.

Robertson, Wright, Dykstra (1988) *Order Restricted Statistical Inference*. Wiley.